THE LANCET

Supplementary webappendix

This webappendix formed part of the original submission and has been peer reviewed. We post it as supplied by the authors.

Supplement to: Smith KR, Jerrett M, Anderson HR, et al. Public health benefits of strategies to reduce greenhouse-gas emissions: health implications of short-lived greenhouse pollutants. *Lancet* 2009; published online Nov 25. DOI:10.1016/S0140-6736(09)61716-5.

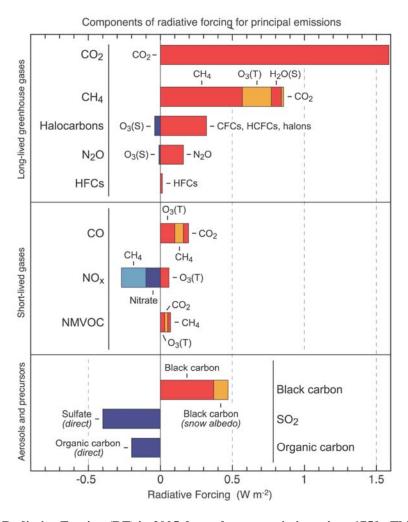
Webappendix

Public Health Benefits of Strategies to Reduce Greenhouse-Gas Emissions: Health Implications of the Short-Lived Greenhouse Pollutants

Table of Contents:

- I. Box Figure 1. Global Warming (Radiative Forcing [RF]) in 2005 due to human emissions since 1750...page 1
- II. Primer on Black Smoke, Black Carbon, and Elemental Carbon Metrics.....page 3
- III. Supplemental Material for Time Series Systematic Reviews.....page 9
- IV. American Cancer Society Cancer Prevention Study II (CPS).....page 26

Box Figure 1. Radiative Forcing (RF) in 2005 due to human emissions since 1750.



Box Figure 1. Radiative Forcing (RF) in 2005 due to human emissions since 1750. This graph is focused on the emitted species and thus shows both the direct and indirect climate impacts that result for each. Several affect ozone levels, for example, but ozone itself is not emitted directly. Some, such as NO_x , have both warming and cooling (negative RF) impacts. Others, like methane, have both direct and indirect impacts. Several are eventually oxidized to CO_2 in the atmosphere and, if from fossil sources, contribute to net warming by this route as well. Based on Forster et al. $(2007)^1$

1. Forster, P., V. Ramaswamy, P. Artaxo, T. Berntsen, R. Betts, D.W. Fahey, J. Haywood, J. Lean, D.C. Lowe, G. Myhre, J. Nganga, R. Prinn, G. Raga, M. Schulz and R. Van Dorland, 2007: Changes in Atmospheric Constituents and in Radiative Forcing. Chapter 2 in: *Climate Change 2007: The Physical Science Basis. Contribution of Working Group I to the Fourth Assessment Report of the Intergovernmental Panel on Climate Change* [Solomon, S., D. Qin, M. Manning, Z. Chen, M. Marquis, K.B. Averyt, M.Tignor and H.L. Miller (eds.)]. Cambridge University Press, Cambridge, United Kingdom and New York, NY, USA.

II. Primer on Black Smoke, Black Carbon, and Elemental Carbon Metrics

Summary

Black carbon (BC) is the term used by the climate-modeling community to describe small soot particles that are highly warming per unit mass. As yet, there is only limited monitoring data of BC for epidemiologic analyses. Black smoke (BS) measurements are closely related, date back to the 1920s, and continue to be collected today. BS data provide the largest evidence base for epidemiologic studies of a measure related to BC. It is not straightforward, however, to convert one to the other in a quantitative way across studies ^{1, 2}. Elemental carbon (EC), determined by a thermal-chemical technique, has a more straightforward relationship with BC and has been measured in a few air pollution monitoring networks.

Measurement methods

The Black Smoke Index

BS measurements, reported in units of µg m⁻³, are taken by pumping a known volume of air through a known area of filter paper, and then measuring the darkness of the stain by its reflectance of white light². The reflectance is measured relative to the light reflected by a clean filter of the same material. These measurements may be made using an instrument such as the EEL M43D Smokestain Reflectometer where the reflectance, volume, and area are converted to a BS Index using a standard table.² The OECD convention for BS is derived from the following equation:

$$I_{\rm BS} (\mu \rm g \, m^{-3}) = 8.655 \times 10^8 \alpha^2 + 2.219 \times 10^5 \alpha$$
 (eq.1)²

where I_{BS} is the BS Index and α is the absorption coefficient of the sampled air, in units of m⁻¹. The British Smoke Stain convention for BS is equal to the OECD version multiplied by a factor of 0.8667. For a complete discussion of the derivation of the Black Smoke Index, please refer to Quincey et al. (2007).²

Black Carbon

The measurement of BC is performed by an optical <u>transmission</u> method, as distinct from the optical <u>reflectance</u> method for BS. The sample is collected on a filter and the optical transmission is measured either in 'batch mode' for a time-weighted average determination; or incrementally while the filter is collecting, to provide real-time data. This may be performed at multiple optical wavelengths to provide additional information or speciation. Following Quincey (2007)², the BC concentration can be determined by the following equation:

$$C_{\rm BC} (\mu g \, {\rm m}^{-3}) = 10^6 \alpha ({\rm m}^{-1}) / \alpha_{\rm atn} ({\rm m}^2 \, {\rm g}^{-1}).$$
 (eq. 2)²

where C_{BC} is the BC concentration; α is the absorption coefficient; and σ_{atn} is the mass extinction coefficient for particles collected on the filter. This, in practice, depends on the particle size, particle composition, wavelength of light used, and filter type; and is determined empirically.² Quincey (2007) uses a σ_{atn} value of 16.6 m²g⁻¹ taken from the Magee AE21 Aethalometer instrument literature. From this,

$$C_{\rm BC} (\mu \rm g \, m^{-3}) = 6.02 \times 10^4 \alpha$$
 _{(eq. 3)²}

For a complete discussion of the derivation of the above equation, refer to Quincey (2007).²

Elemental Carbon

The material denoted "Elemental" Carbon (EC) is determined by thermal-optical analysis in a multi-step process. The particulate matter (PM) is usually collected on quartz fiber filters and then analyzed at

different temperatures in inert or oxidizing atmospheres. Carbon released from the sample is converted to CO_2 or methane and detected during the progressive analysis. Assignment of the precise fraction of the analyte to "EC" is still an active topic of research^{3,4}, and there are several different protocols, see Environmental Protection Agency (2003)⁵. These include the National Institute of Occupational Safety and Health (NIOSH) method 5040; the "IMPROVE-A" protocol developed by the Desert Research Institute, NV; and the "EUSAAR" protocol of JRC, Ispra (Europe). In this paper, we used EC data from ambient air PM samples analyzed by the NIOSH method.

Because there are many techniques that have been employed to estimate EC concentrations, a number of studies have examined inter-method comparability. As an example, Miller et al. (2007)⁶ compared the NIOSH method with the Thermo Electron Personal Data RAM 1200 photometer (PDR) technique (particulate mass concentration based on light scattering) and found that they agreed well (R²=0.97) when the data were adjusted for relative humidity (Figure 1).⁶

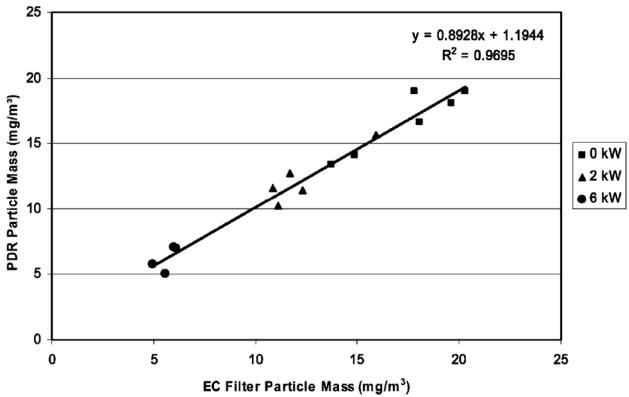


Figure 1: Linear regression of corrected PDR data and EC filter (NIOSH) sample data. Adapted From Miller et al. (2007).

Translation between metrics

Black Smoke to Black Carbon

Because both BS and BC methods essentially measure the same thing – the optical absorption coefficient of airborne particles collected on a filter – Quincey (2007) has shown that a simple, quadratic equation can effectively convert black smoke measurements into black carbon concentration estimates²:

$$I_{\rm BS} = 0.239C_{\rm BC}^2 + 3.69C_{\rm BC}$$
 (eq. 4)²

or, conversely,

$$C_{\rm BC} = \sqrt{4.18I_{\rm BS} + 59.6} - 7.72$$
(eq. 5)²

From another, albeit more simple methodological angle, Novakov and Hansen $(2004)^7$ follow Erdman et al. $(1993)^8$ and convert their BS data, in units of μg (BS) m^{-3} , to BC data in units of μg (BC) m^{-3} via the linear relationship, BC = 0.23 BS.

The relationships between the BS index and BC (μ g m⁻³) are robust as shown by the tight scatters (Figure 2) from the modeled data from Quincey et al. (2007); and Figure 4 shows this relationship with high R² values (urban=0.85; rural=0.75) for empirical data from the Netherlands.⁹

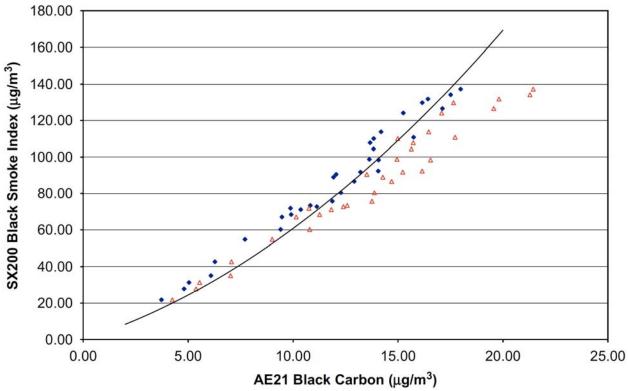


Figure 2: Unprocessed Black Carbon data (dark diamonds) and processed Black Carbon data (light triangles) plotted against the SX200 Black Smoke data, together with the derived theoretical relationship between Black Carbon and Black Smoke measurements (Eq. (4) – black line). From Quincey et al. (2007)²

Elemental Carbon to Black Carbon

Although subject to the measurement complexities described above, there is a high correlation between EC and BC (R^2 =0.89) (Figure 3). ¹² Both optical and thermal methods can account for, or separate, the contributions from 'graphitic' carbon or 'tarry' organic compounds²: optical methods can analyze at multiple wavelengths while thermal methods can employ different temperature profiles or oxidizing conditions. EC is interchangeable with BC insofar as both of the components, regardless of their physiochemical definition, represent the vast majority of light-absorbing components in undifferentiated PM and are thus significant warming agents. ¹¹ Because of this, many researchers struggle with criteria that can distinguish between the two and often conflate the terms. ¹³

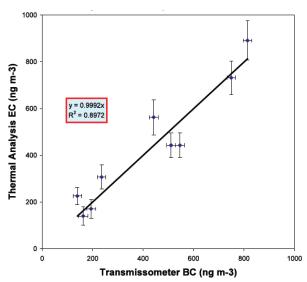


Figure 3: Comparison of Optical Transmission (BC) measurements with Thermal (EC) Analysis. Data Source: Ahmeda et al. (2009). 12

Moreover, Schaap et al. (2007)⁹ notes that the relation between BS and EC is linear and highly correlated, albeit dependent on the type of location (Figure 4). The authors propose that EC can be predicted using the following relationships for rural and urban locations:

$$EC = \begin{cases} 0.056 \times BS + 0.16 & (rural), \\ 0.088 \times BS + 0.32 & (urban). \end{cases}$$
 (eq. 6) Cited from: Schaap et al. (2007)⁹

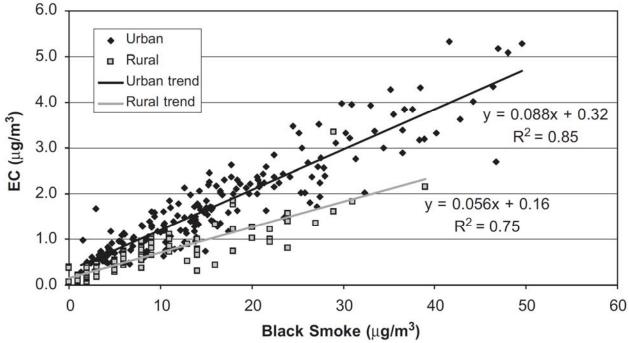


Figure 4: Elemental carbon as a function of Black Smoke at rural and urban locations. Cited from: Schaap et al. (2007)⁹

There are a wide range of inter-species delineations between EC and BC, as well as intra-species differences between various physiochemical forms of BC. Both EC and BC data can be used for

environmental epidemiology for the following reasons:

- 1. EC and BC data both represent emissions from combustion processes. Neither represents a unique chemical formulation.
- Both BC and EC data metrics indicate the mass of nearly all the light-absorbing constituents of undifferentiated PM.¹¹
- 3. Carbonaceous particles of combustion origin tend to be diametrically small¹⁴⁻¹⁸. PM of diesel origin tends to be finer than that from other sources such as coal or biomass combustion. ¹⁹ Ultrafine particles, however, tend to agglomerate soon after release from the emission source²⁰, and BC is usually found in the ambient atmosphere in the size ranges of 0.1 to 0.5 μm. Nonetheless, it is thought that EC and BC may have effects both chemically and as a generic function of being particles (see main paper for further discussion).
- 4. These carbonaceous particles are nearly always associated with other often toxic constituents such as polycyclic aromatic hydrocarbons (PAHs), and other organic compounds.²⁰ The California Air Resources Board (CARB) uses BC as a surrogate for Diesel Particulate Matter (DPM), which has been designated as a Toxic Air Contaminant (TAC)²¹

As the cohort we examine in the main paper is all urban and in one country, we consider that EC is as good an indicator of what climate scientists refer to as BC as can currently be found in air pollution measurement networks.

Sources Cited - Primer on Black Smoke, Black Carbon, and Elemental Carbon Metrics

- Heal MR, Hibbs LR, Agius RM, Beverland LJ. Interpretation of variations in fine, coarse and black smoke particulate matter concentrations in a northern European city. Atmospheric Environment. 2005 Jun;39(20):3711-8.
- 2. Quincey P. A relationship between Black Smoke Index and Black Carbon concentration. Atmospheric Environment. 2007 Nov;41(36):7964-8.
- 3. Bae MS, Schauer JJ, Turner JR, Hopke PK. Seasonal variations of elemental carbon in urban aerosols as measured by two common thermal-optical carbon methods. Sci Total Environ. 2009 Sep;407(18):5176-83.
- CDC. NIOSH Method 5040, Elemental Carbon (Diesel Particulate): NIOSH Manual of Analytical Methods, 4th ed. Atlanta, GA: National Institute of Occupational and Environmental Safety; 1997 January 15
- 5. EPA. Standard operating procedure for the determination of organic, elemental, and total carbon in particulate matter using a thermal/optical-transmittance carbon analyzer. Washington, DC: OC/EC Laboratory, Environmental and Industrial Science Division, Research Triangle Institute; Downloaded from the United States Environmental Protection Agency; 2003.
- 6. Miller AL, Habjan MC, Park K. Real-time estimation of elemental carbon emitted from a diesel engine. Environmental Science & Technology. 2007 Aug;41(16):5783-8.
- 7. Novakov T, Hansen JE. Black carbon emissions in the United Kingdom during the past four decades: an empirical analysis. Atmospheric Environment. 2004 Aug;38(25):4155-63.
- 8. Erdman A, Israel G, Ulrich E. Comparative measurements of atmospheric elemental carbon using the British Black Smoke sampler and a thermal carbon analyzer. Staub. 1993;53:187-91.
- 9. Schaap M, van der Gon H. On the variability of Black Smoke and carbonaceous aerosols in the Netherlands. Atmospheric Environment. 2007 Sep;41(28):5908-20.
- 10. Bailey DLR, Clayton P. The measurement of suspended particle and total carbon concentrations in the atmosphere using standard smoke shade methods. Atmospheric Environment. 1982;16(11):2683-90.
- 11. Ramanathan V, Ramana MV, Roberts G, Kim D, Corrigan C, Chung C, et al. Warming trends in Asia amplified by brown cloud solar absorption. Nature. 2007;448(7153):575-U5.
- 12. Ahmeda T, Dutkiewicz V, Shareef A, Tuncel G, Tuncel C, Husain L: Measurement of black carbon (BC) by an optical method and a thermal-optical method: Intercomparison for four sites. In Press. Atmospheric Environment (2009), doi:10.1016/j.atmosenv.2009.09.031
- 13. Ban-Weiss GA, Lunden MM, Kirchstetter TW, Harley RA. Measurement of Black Carbon and Particle Number Emission Factors from Individual Heavy-Duty Trucks. Environmental Science & Technology. 2009 Mar;43(5):1419-24.

- 14. Kamboures MA, Raff JD, Miller Y, Phillips LF, Finlayson-Pitts BJ, Gerber RB. Complexes of HNO3 and NO3 with NO2 and N2O4, and their potential role in atmospheric HONO formation. Phys Chem Chem Phys. 2008 Oct 21;10(39):6019-32.
- 15. Kleeman MJ, Cass GR, Eldering A. Modeling the airborne particle complex as a source-oriented external mixture. Journal of Geophysical Research-Atmospheres. 1997 Sep;102(D17):21355-72.
- 16. Kleeman MJ, Schauer JJ, Cass GR. Size and composition distribution of fine particulate matter emitted from motor vehicles. Environmental Science & Technology. 2000 Apr;34(7):1132-42.
- 17. Ramanathan V, Crutzen PJ, Lelieveld J, Mitra AP, Althausen D, Anderson J, et al. Indian Ocean Experiment: An integrated analysis of the climate forcing and effects of the great Indo-Asian haze. Journal of Geophysical Research-Atmospheres. 2001 Nov;106(D22):28371-98.
- 18. WHO. Health relevance of particulate matter from various sources: report on a WHO workshop, Bonn, Germany, March 2007. Copenhagen: World Health Organization, Regional Office for Europe; 2007.
- 19. Kleeman MJ, Schauer JJ, Cass GR. Size and composition distribution of fine particulate matter emitted from motor vehicles. Environmental Science & Technology. 2000 Apr;34(7):1132-42.
- 20. Moffet R, Prather K. In-situ measurements of the mixing state and optical properties of soot with implications for radiative forcing estimates. Proc Natl Acad Sci U S A. 2009;106(29):11872–7.
- 21. CARB. Proposed identification of diesel exhaust as a Toxic Air Contaminant. Appendix III, Part A, Exposure Assessment. Sacramento, CA: California Air Resources Board; 1998.

III. Supplemental material for Time Series Systematic Reviews

Methods

The Air Pollution Epidemiology Database (APED), a database of results from time series studies of the health effects of outdoor air pollution, provided city specific standardized effect estimates for sulphate (SO₄²) and BS. Details of APED have been published elsewhere 1,2. In brief, the database comprises a bibliography of time series studies published up to May 2009 and a relational database of effect estimates and associated information extracted from selected studies. Published, peer reviewed time series studies were identified through systematic searching of Medline, Embase, and Web of Science databases with no constraints placed on language or date of publication. The search criteria were developed using known papers and then tested several times against citations in published reviews of the literature. A two stage sifting and assessment process was then applied to identify daily time series studies providing the necessary data for quantitative analysis. This assessment considered the appropriateness of the statistical methods used, data quality and the provision of the information needed to standardize the effect estimates. The data characterizing the effect estimates were entered into a relational database (ACCESS, Microsoft Corp.) where routines to calculate standardized estimates were applied. Studies reporting associations between all cause, cardiovascular and respiratory mortality and SO_4^{-2} , BS and O_3 were identified and their results extracted into STATA. City specific results were included in a meta-analysis if the estimate was accompanied by an indication of the precision of the estimate. When more than one estimate for a specific location was available only the most recently published estimate for that location was included in the metaanalysis. Summary fixed and random effects estimates were calculated using the method of Der Simonian and Laird³ and implemented in the *metan* procedure in STATA⁴. Results from multi- city studies were also extracted and although not included in the calculation of summary estimates their inclusion in this review was warranted because of the additional evidence they provide.

Sources Cited: Supplemental material for time series systematic reviews – Methods

- Anderson HR, Atkinson WA, Peacock JL, Marston L, Konstantinou K. Meta-analysis of time-series studies and panel studies of particulate matter (PM) and ozone (O3). Report of a WHO task group. World Health Organization 2004. http://www.euro.who.int/document/E82792.pdf (last accessed on 09.07.2009)
- 2. Anderson HR et al. Ambient Particulate Matter and Health Effects: Publication Bias in Studies of Short-Term Associations Epidemiol 2005;16:155-163
- 3. Der Simonian R, Laird N. Meta-analysis in clinical trials. Controlled Clinical Trials.1986;7:177-188
- 4. Bradburn MJ, Deeks JJ, Altman DG. sbe24: metan an alternative meta-analysis command. Stata Technical Bulletin 1998; 44: 4-15.

Table 1 City specific, fixed and random effects summary estimates and multi-city estimates for sulphate and all cause, all age mortality Single city/site estimates

Author	Year	Country	City	Conc*	%°	95% CI ^{\$}
Dockery	1992	USA	St. Louis	8.7	0.61	(-0.52, 1.75)
Dockery	1992	USA	Tennessee eastern	8.0	0.80	(-1.54, 3.20)
Burnett	1998	Canada	Toronto	8.2	0.18	(0.08, 0.27)
Gwynn	2000	USA	Buffalo	5.9	0.23	(0.05, 0.41)
Klemm	2000	USA	Georgia	4.4	0.51	(-0.54, 1.57)
Hoek	2000	Netherlands	Netherlands	3.8	0.13	(0.02, 0.23)
Lippmann	2000	USA	Wayne County	3.2	0.15	(-0.12, 0.42)
Goldberg	2001	Canada	Montreal	2.2	0.51	(0.27, 0.76)
Anderson	2001	UK	West Midlands	2.7	-0.07	(-0.40, 0.26)
Maynard	2007	USA	Boston	2.4	0.49	(0.00, 0.97)
			Pooled FE		0.18	(0.12, 0.24)
			Pooled RE		0.21	(0.11, 0.30)
			Heterogeneity chi-squar	red = 13.55 (e	d.f. = 9) p	= 0.139

Multi-city estimates

Author	Year	Location	City	%°	95% CI ^{\$}
Schwartz	1996	USA	6 communities	0.22	(0.13, 0.31)
Brook	2007	Canada	10 cities	0.40	(0.10, 0.70)
Ostro	2007	USA	6 California counties	0.12	(-0.76, 1.00)

Table 2 City specific, fixed and random effects summary estimates and multi-city estimates for sulphate and all cardiovascular/cardiac, all age mortality

^{*} Mean/median daily concentrations (μ g/m3) \sim Percentage change in mean number of deaths per 1μ g/m3 increase in SO_4^{-2}

^{\$ 95%} confidence limits

Single city estimates

Author	Year	Country	City	Conc*	%~	95%CI~
Gwynn	2000	USA	Buffalo	5.9	0.26	(-0.03, 0.55)
Hoek	2000	Netherlands	Netherlands	3.8	0.08	(-0.08, 0.24)
Lippmann	2000	USA	Wayne County	3.2	0.11	(-0.29, 0.52)
Anderson	2001	UK	West Midlands	2.7	-0.23	(-0.70, 0.25)
Maynard	2007	USA	Boston	2.4	-0.09	(-0.67, 0.49)
			Pooled FE		0.09	(-0.04, 0.21)
			Pooled RE		0.09	(-0.04, 0.21)
			Heterogeneity chi-square	d = 3.32 (d.	f. = 4) p = 0.5	506

Multi-city estimates

Author	Year	Location	City	% ~~	95%CI~	
Ostro	2007	USA	6 California counties	0.36	(-0.91, 1.65)	

Table 3 City specific, fixed and random effects summary estimates and multi-city estimates for sulphate and all respiratory, all age mortality

Single city estimates

Author	Year	Country	City	Conc*	%°	95% CI ^{\$}
Gwynn	2000	USA	Buffalo	5.9	0.68	(0.05, 1.31)
Lippmann	2000	USA	Wayne County	3.2	0.40	(-0.61, 1.43)
Anderson	2001	UK	West Midlands	2.7	-0.31	(-1.13, 0.52)
Maynard	2007	USA	Boston	2.4	0.92	(-0.49, 2.36)
			Pooled FE		0.39	(-0.04, 0.82)
			Pooled RE		0.37	(-0.15, 0.90)
			Heterogeneity chi-squar	red = 4.11	(d.f. = 3) p	=0.250

^{*} Mean/median daily concentrations ($\mu g/m3$) \sim Percentage change in mean number of deaths per $1\mu g/m3$ increase in SO42- \$95% confidence limits

Multi-city estimates

Author	Year	Location	City	0 ∕₀~	95% CI ^{\$}
Ostro	2007	USA	6 California counties	0.70	(-1.65, 3.11)

^{*} Mean/median daily concentrations ($\mu g/m3$) \sim Percentage change in mean number of deaths per $1\mu g/m3$ increase in SO42- \$95% confidence limits

Table 4 City specific, fixed and random effects summary estimates and multi-city estimates for Black Smoke and all cause, all age mortality Single city estimates

Author	Year	Country	City	Conc*	%°	95% CI ^{\$}
Bellido Blasco	1999	Spain	Castellon	20.3	0.15	(-0.05, 0.35)
Aguinaga	1999	Spain	Pamplona	21.7	0.29	(-0.19, 0.78)
Perez	1999	Spain	Vitoria-Gasteiz	45.0	0.06	(-0.03, 0.16)
Katsouyanni	2001	Greece	Athens	64.0	0.07	(0.04, 0.09)
Katsouyanni	2001	Spain	Barcelona	39.0	0.16	(0.10, 0.21)
Katsouyanni	2001	Spain	Bilbao	23.0	0.08	(-0.07, 0.23)
Katsouyanni	2001	UK	Birmingham	11.0	0.03	(-0.06, 0.13)
Katsouyanni	2001	Poland	Cracow	36.0	-0.02	(-0.06, 0.02)
Katsouyanni	2001	Slovenia	Ljubljana	13.0	-0.01	(-0.13, 0.11)
Katsouyanni	2001	Poland	Lodz	30.0	-0.01	(-0.05, 0.04)
Katsouyanni	2001	Poland	Poznan	23.0	0.06	(0.02, 0.11)
Saez	2001	Spain	Vigo	102.6	0.07	(0.01, 0.13)
Anderson	2001	UK	West Midlands	10.9	0.04	(-0.09, 0.16)
Katsouyanni	2001	Poland	Wroclaw	33.0	0.03	(-0.02, 0.07)
Arribas-Monzon	2001	Spain	Zaragoza	n/a	0.04	(-0.04, 0.11)
Le Tertre	2002	France	Bordeaux	13.2	0.15	(0.00, 0.30)
Le Tertre	2002	France	Le Havre	n/a	0.02	(-0.15, 0.20)
Le Tertre	2002	France	Marseille	16.0	0.12	(0.05, 0.19)
Le Tertre	2002	France	Paris	16.0	0.04	(0.00, 0.08)
Le Tertre	2002	France	Rouen	14.1	0.01	(-0.10, 0.13)
Ballester	2002	Spain	Valencia	44.2	0.18	(0.09, 0.27)
Pattenden	2003	UK	London	n/a	0.08	(-0.67, 0.84)
Goodman	2004	Ireland	Dublin	40.0	0.04	(0.03, 0.05)
Bogdanovic	2006	Serbia	Nis	23.0	0.11	(0.01, 0.22)
Fischer	2009	Netherlands	Netherlands	6.0	-0.03	(-0.10, 0.04)

 Pooled FE
 0.04 (0.04, 0.05)

 Pooled RE
 0.05 (0.03, 0.07)

 Heterogeneity chi-squared =
 62.09 (d.f. = 24) p = 0.000

Multi-city estimates

Author	Year	Location	City	%	95% CI ^{\$}
Katsouyanni	1997	Europe	APHEA 1	0.03	(0.02, 0.03)
Aga	2003	Europe	APHEA 2 - 29 cities	0.06	(0.03, 0.08)
Ballester	2003	Spain	13 cities	0.08	(0.04, 0.12)
Carder	2008	Scotland	9 sites	0.17	(0.07, 0.26)

^{*} Mean/median daily concentrations ($\mu g/m3$) ~ Percentage change in mean number of deaths per 1 $\mu g/m3$ increase in BS \$ 95% confidence limits

Table 5 City specific, fixed and random effects summary estimates and multi-city estimates for Black Smoke and all cardiovascular/cardiac, all age mortality

Single city estimates

Author	Year	Country	City	Conc*	%°	95% CI ^{\$}
Wojtyniak	1996	Poland	Krakow	73.3	0.01	(-0.02, 0.05)
Wojtyniak	1996	Poland	Lodz	57.3	0.01	(-0.02, 0.05)
Wojtyniak	1996	Poland	Poznan	34.0	-0.02	(-0.08, 0.04)
Wojtyniak	1996	Poland	Wroclaw	54.3	0.01	(-0.04, 0.06)
Cambra	1999	Spain	Bilbao	23.1	-0.17	(-0.37, 0.04)
Bellido Blasco	1999	Spain	Castellon	20.3	0.34	(0.05, 0.64)
Bremner	1999	UK	London	10.8	0.12	(-0.01, 0.25)
Aguinaga	1999	Spain	Pamplona	21.7	-0.23	(-1.04, 0.58)
Garcia-		•	•			
Aymerich	2000	Spain	Barcelona	42.4	0.11	(0.04, 0.19)
Anderson	2001	ÚK	West Midlands	10.9	0.09	(-0.09, 0.27)
Arribas-						
Monzon	2001	Spain	Zaragoza	n/a	0.07	(-0.05, 0.18)
Le Tertre	2002	France	Bordeaux	13.2	0.17	(-0.07, 0.42)
Le Tertre	2002	France	Marseille	16.0	0.09	(-0.02, 0.21)
Le Tertre	2002	France	Paris	16.0	0.04	(-0.02, 0.10)
Le Tertre	2002	France	Rouen	14.1	0.10	(-0.11, 0.30)
Ballester	2002	Spain	Valencia	44.2	0.15	(0.01, 0.29)
Goodman	2004	Ireland	Dublin	40.0	0.04	(0.02, 0.06)
Filleul	2006	France	Le Havre	11.0	0.44	(-0.09, 0.97)
Stankovic	2007	Serbia	Nis	22.8	0.13	(-0.03, 0.29)
Fischer	2009	Netherlands	Netherlands	6.0	-0.12	(-0.24, 0.00)
			n i ier		0.02	(0.02.0.04)
			Pooled FE		0.03	(0.02, 0.04)
			Pooled RE	1 0	0.04	(0.01, 0.06)
			Heterogeneity chi	-squared = 36	5.82 (d.f. =	19) $p = 0.008$

Multi-city estimates

Author	Year	Location	City	%°	95% CI ^{\$}
Analitis	2006	Europe	APHEA 2 - 15 cities	0.06	(0.03, 0.09)
Ballester	2003	Spain	13 cities	0.03	(-0.02, 0.08)
Carder	2008	Scotland	9 sites	0.04	(-0.10, 0.18)

^{*} Mean/median daily concentrations (μ g/m3) \sim Percentage change in mean number of deaths per 1 μ g/m3 increase in BS \$ 95% confidence limits

Table 6 City specific, fixed and random effects summary estimates and multi-city estimates for Black Smoke and all respiratory, all age mortality Single city estimates

Author	Year	Country	City	Conc*	%°	95% CI ^{\$}
Wojtyniak	1996	Poland	Krakow	73.3	-0.02	(-0.14, 0.09)
Wojtyniak	1996	Poland	Lodz	57.3	-0.08	(-0.19, 0.02)
Wojtyniak	1996	Poland	Poznan	34.0	-0.10	(-0.28, 0.08)
Wojtyniak	1996	Poland	Wroclaw	54.3	-0.19	(-0.34, -0.03)
Cambra	1999	Spain	Bilbao	23.1	0.29	(-0.11, 0.70)
Bellido Blasco	1999	Spain	Castellon	20.3	0.36	(-0.26, 0.98)
Bremner	1999	UK	London	10.8	0.19	(0.02, 0.36)
Aguinaga	1999	Spain	Pamplona	21.7	1.26	(-0.39, 2.94)
Garcia-Aymerich	2000	Spain	Barcelona	42.4	0.10	(-0.05, 0.25)
Anderson	2001	UK	West Midlands	10.9	0.01	(-0.29, 0.31)
Arribas-Monzon	2001	Spain	Zaragoza	n/a	0.29	(0.06, 0.51)
Le Tertre	2002	France	Bordeaux	13.2	0.20	(-0.37, 0.77)
Le Tertre	2002	France	Marseille	16.0	0.26	(0.02, 0.50)
Le Tertre	2002	France	Paris	16.0	-0.02	(-0.15, 0.11)
Le Tertre	2002	France	Rouen	14.1	0.07	(-0.33, 0.46)
Ballester	2002	Spain	Valencia	44.2	-0.21	(-0.50, 0.08)
Goodman	2004	Ireland	Dublin	40.0	0.09	(0.05, 0.13)
Filleul	2006	France	Le Havre	11.0	0.45	(-0.64, 1.55)
Bogdanovic	2006	Serbia	Nis	23.0	0.17	(-0.09, 0.44)
Fischer	2009	Netherlands	Netherlands	6.0	0.03	(-0.19, 0.25)

 Pooled FE
 0.06
 (0.03, 0.09)

 Pooled RE
 0.04
 (-0.02, 0.11)

Heterogeneity chi-squared = 43.88 (d.f. = 19) p = 0.001

Multi-city estimates

Author	Year	Location	City	%°	95% CI ^{\$}
Analitis	2006	Europe	APHEA 2 - 15 cities	0.08	(0.01, 0.16)
Ballester	2003	Spain	13 cities	0.11	(0.04, 0.18)
Carder	2008	Scotland	9 sites	0.52	(0.29, 0.76)

Table 7 City specific, fixed and random effects summary estimates and multi-city estimates for ozone (8 hour) and all cause, all age mortality

Single city estimates

Author	Year	Country	City	Conc*	%°	95% CI ^{\$}
Simpson	1997	Australia	Brisbane	16.7	0.12	(0.04, 0.20)
Borja-Aburto	1997	Mexico	Mexico			(0.01, 0.03)
v			City	94.0	0.02	
Michelozzi	1998	Italy	Rome	21.0	0.04	(0.00, 0.08)
Hong	1999	South Korea	Inchon	15.4	-0.02	(-0.05, 0.00)
Bremner	1999	UK	London	16.0	-0.01	(-0.05, 0.02)
Cadum	1999	Italy	Turin	73.7	0.03	(-0.01, 0.08)
Klemm	2000	USA	Georgia	40.6	-0.02	(-0.13, 0.09)
Hoek	2000	Netherlands	Netherlands	47.0	0.02	(0.01, 0.03)
Wong	2001	China	Hong Kong	33.5	0.02	(-0.02, 0.05)
Anderson	2001	UK	West			(0.00, 0.10)
			Midlands	24.0	0.05	
Saez	2002	Spain	Barcelona	67.5	0.01	(-0.02, 0.04)
Le Tertre	2002	France	Le Havre	43.4	0.07	(-0.05, 0.18)
Le Tertre	2002	France	Lyon	52.0	0.07	(0.00, 0.15)
Saez	2002	Spain	Madrid	42.1	0.03	(-0.01, 0.06)
Le Tertre	2002	France	Paris	26.0	0.04	(0.00, 0.08)
Le Tertre	2002	France	Rouen	57.9	0.10	(0.00, 0.21)

^{*} Mean/median daily concentrations (ug/m3) \sim Percentage change in mean number of deaths per 1 µg/m3 increase in BS \$95% confidence limits

Le Tertre	2002	France	Strasbourg	37.0	0.06	(-0.01, 0.12)	
Le Tertre	2002	France	Toulouse	68.0	0.03	(-0.08, 0.14)	
Saez	2002	Spain	Valencia	45.5	0.22	(0.05, 0.40)	
Parodi	2005	Italy	Genoa	79.2	0.06	(0.01, 0.11)	
Zhang	2006	China	Shanghai	56.1	0.04	(0.02, 0.07)	
Qian	2007	China	Wuhan	47.2	0.02	(-0.01, 0.05)	
			Pooled FE		0.02	(0.02, 0.03)	
			Pooled RE		0.03	(0.02, 0.04)	
			Heterogeneity chi-squared = 47.55 (d.f. = 21) p = 0.001				

Multi-city estimates

Author	Year	Location	City	Conc*	%°	95% CI ^{\$}
Gryparis	2004	Europe	23 cities	n/a	0.003	(-0.018, 0.024)

Table 8 City specific, fixed and random effects summary estimates and multi-city estimates for ozone (8 hour) and cardiovascular, all age mortality Single city estimates

Year	Country	City	Conc*	%~ ~	95% CI ^{\$}
1997	Mexico	Mexico			(0.02, 0.06)
		City	94.0	0.04	
1999	South Korea	Inchon	15.4	-0.09	(-0.65, 0.47)
1999	UK	London	16.0	0.07	(0.01, 0.12)
1999	Italy	Turin	73.7	0.07	(0.00, 0.14)
2001	Netherlands	Netherlands	n/a	0.04	(0.02, 0.05)
2001	UK	West			(-0.06, 0.09)
		Midlands	24.0	0.02	
2002	Spain	Barcelona	67.5	0.06	(0.00, 0.11)
	1997 1999 1999 1999 2001 2001	1997 Mexico 1999 South Korea 1999 UK 1999 Italy 2001 Netherlands 2001 UK	1997 Mexico Mexico City 1999 South Korea Inchon 1999 UK London 1999 Italy Turin 2001 Netherlands Netherlands 2001 UK West Midlands	1997 Mexico Mexico City 94.0 1999 South Korea Inchon 15.4 1999 UK London 16.0 1999 Italy Turin 73.7 2001 Netherlands Netherlands n/a 2001 UK West Midlands 24.0	1997 Mexico Mexico City 94.0 0.04 1999 South Korea Inchon 15.4 -0.09 1999 UK London 16.0 0.07 1999 Italy Turin 73.7 0.07 2001 Netherlands Netherlands n/a 0.04 2001 UK West Midlands 24.0 0.02

^{*} Mean/median daily concentrations (ug/m3) ~ Percentage change in mean number of deaths per 1 ug/m3 increase in O3

^{\$ 95%} confidence limits

Wong	2002	China	Hong Kong	29.3	-0.03	(-0.09, 0.03)
Le Tertre	2002	France	Le Havre	43.4	0.03	(-0.18, 0.24)
Le Tertre	2002	France	Lyon	52.0	0.04	(-0.16, 0.24)
Saez	2002	Spain	Madrid	42.1	0.05	(0.00, 0.11)
Le Tertre	2002	France	Paris	26.0	0.04	(-0.03, 0.12)
Le Tertre	2002	France	Rouen	57.9	0.14	(-0.06, 0.34)
Le Tertre	2002	France	Strasbourg	37.0	0.02	(-0.08, 0.13)
Le Tertre	2002	France	Toulouse	68.0	0.10	(-0.09, 0.29)
Saez	2002	Spain	Valencia	45.5	0.31	(0.04, 0.58)
Penttinen	2004	Finland	Helsinki	50.0	0.04	(-0.05, 0.12)
Zhang	2006	China	Shanghai	56.1	0.05	(0.01, 0.10)
Qian	2007	China	Wuhan	47.2	0.02	(-0.05, 0.10)
			Pooled FE		0.04	(0.03, 0.05)
			Pooled RE		0.04	(0.03, 0.05)

Table 9 City specific, fixed and random effects summary estimates and multi-city estimates for ozone (8 hour) and respiratory, all age mortality Single city estimates

Heterogeneity chi-squared = 13.61 (d.f. = 18) p = 0.754

Author	Year	Country	City	Conc*	% ~	95% CI ^{\$}
Simpson	1997	Australia	Brisbane	16.7	0.19	(-0.09, 0.48)
Borja-Aburto	1997	Mexico	Mexico			(-0.02, 0.05)
			City	94.0	0.02	
Hong	1999	South Korea	Inchon	15.4	0.01	(-0.90, 0.92)
Bremner	1999	UK	London	16.0	-0.07	(-0.16, 0.01)
Cadum	1999	Italy	Turin	73.7	0.05	(-0.15, 0.25)
Anderson	2001	UK	West			(-0.10, 0.17)
			Midlands	24.0	0.04	
Saez	2002	Spain	Barcelona	67.5	0.04	(-0.06, 0.14)

^{*} Mean/median daily concentrations (ug/m3) ~ Percentage change in mean number of deaths per 1 ug/m3 increase in O3 \$ 95% confidence limits

Wong	2002	China	Hong			(0.04, 0.16)
			Kong	29.3	0.10	
Le Tertre	2002	France	Le Havre	43.4	-0.20	(-0.67, 0.27)
Le Tertre	2002	France	Lyon	52.0	0.17	(-0.10, 0.44)
Saez	2002	Spain	Madrid	42.1	0.01	(-0.10, 0.12)
Le Tertre	2002	France	Paris	26.0	-0.04	(-0.19, 0.11)
Le Tertre	2002	France	Rouen	57.9	0.21	(-0.21, 0.62)
Le Tertre	2002	France	Strasbourg	37.0	0.00	(-0.24, 0.25)
Le Tertre	2002	France	Toulouse	68.0	0.10	(-0.32, 0.52)
Saez	2002	Spain	Valencia	45.5	0.18	(-0.39, 0.76)
Penttinen	2004	Finland	Helsinki	50.0	0.32	(0.12, 0.52)
Zhang	2006	China	Shanghai	56.1	0.03	(-0.04, 0.11)
Qian	2007	China	Wuhan	47.2	0.06	(-0.04, 0.17)

 $Pooled\ FE$ 0.03 (0.01, 0.06) Pooled RE 0.04 (0.01, 0.07) Heterogeneity chi-squared = 25.46 (d.f. = 18) p = 0.113

^{*} Mean/median daily concentrations (ug/m3) ~ Percentage change in mean number of deaths per 1 ug/m3 increase in O3 \$ 95% confidence limit

Sources Cited: Time Series Systematic Reviews

Aga E, Samoli E, Touloumi G, Anderson HR, Cadum E, Forsberg B, Goodman P et al. Short-term effects of ambient particles on mortality in the elderly: results from 28 cities in the APHEA2 project. European Respiratory Journal 2003; 21: S28-S33.

Analitis A, Katsouyanni K, Dimakopoulou K, Samoli E, Nikoloulopoulos AK, Petasakis Y, Touloumi G, Schwartz J, Anderson HR, Cambra K, Forastiere F, Zmirou D, Vonk JM, Clancy L, Kriz B, Bobvos J, Pekkanen J. Short-Term Effects of Ambient Particles on Cardiovascular and Respiratory Mortality. Epidemiology 2006; 17 (2):230-33

Anderson HR, Bremner SA, Atkinson RW, Harrison RM, Walters S. Particulate matter and daily mortality and hospital admissions in the west midlands conurbation of the United Kingdom: associations with fine and coarse particles, black smoke and sulphate. Occup Environ Med 2001;58(8):504-10.

Aguinaga OI, Guillen GF, Oviedo de Sola PJ, Floristan Floristan MY, Laborda Santesteban MS, Martinez Ramirez MT et al. The short-term effects of air pollution on mortality: the results of the EMECAM project in the city of Pamplona, 1991-95. Estudio Multicentrico Espanol sobre la Relacion entre la Contaminacion Atmosferica y la Mortalidad. [Spanish]. Rev Esp Salud Publica 1999;73(2):253-8.

Arribas-Monzon F, Rabanaque MJ, Martos MC, Abad JM, Alcala-Nalvaiz T, Navarro-Elipe M. Effects of air pollution on daily mortality in Zaragoza, Spain, 1991-1995. Salud Publica Mex 2001;43(4):289-97.

Ballester F, Iniguez C, Perez-Hoyos S, Tenias JM. Particulate air pollution and health in Valencia. Gaceta Sanitaria 2002 November; 16(6):464-79.

Ballester F, Iniguez C, Saez M, Perez-Hoyos S, Daponte A, Ordonez JM, Barcelo MA. Short-term relationship between air pollution and mortality in 13 Spanish cities. Medicina Clinica 2003;121(18): 684-689.

Bellido Blasco JB, Daudi CF, Arnedo PA, Gonzalez MF, Herrero CC, Safont AL. The short-term effects of air pollution on mortality: the results of the EMECAM project in Castellon, 1991-95. Estudio Multicentrico Espanol sobre la Relacion entre la Contaminacion Atmosferica y la Mortalidad. [Spanish]. Rev Esp Salud Publica 1999;73(2):225-31.

Bogdanovic D, Nikic D, Milosevic Z, Stankovic A. Black smoke air pollution and daily non-accidental mortality in Nis, Serbia. Central European Journal of Medicine 2006;1(3):292–297.

Borja-Aburto VH, Loomis DP, Bangdiwala SI, Shy CM, Rascon-Pacheco RA. Ozone, suspended particulates, and daily mortality in Mexico City. American Journal of Epidemiology 1997;145(3):258-268.

Bremner SA, Anderson HR, Atkinson RW, McMichael AJ, Strachan DP, Bland JM et al. Short-term associations between outdoor air pollution and mortality in London 1992-4. Occup Environ Med 1999;56(4):237-44.

Brook JR, Burnett RT, Dann TF, Cakmak S, Goldberg MS, Fan XH, Wheeler AJ. Further interpretation of the acute effect of nitrogen dioxide observed in Canadian time-series studies. Journal of Exposure Science and Environmental Epidemiology 2007;17: S36–S44.

Burnett RT, Cakmak S, Raizenne ME, Stieb D, Vincent R, Krewski D et al. The association between ambient carbon monoxide levels and daily mortality in Toronto, Canada. J Air Waste Manage Assoc 1998;48(8):689-700.

Cadum E, Rossi G, Mirabelli D, Vigotti MA, Natale P, Albano L, Marchi G et al. Air pollution and daily mortality in Turin, 1991-1996. Epidemiologia e Prevenzione 1999;23(4):268-276.

Cambra CK, Alonso FE. The short-term effects of air pollution on mortality: the results of the EMECAM project in greater Bilbao. Estudio Multicentrico Espanol sobre la Relacion entre la Contaminacion Atmosferica y la Mortalidad. [Spanish]. Rev Esp Salud Publica 1999;73(2):209-14.

Carder M, McNamee R, Beverland I, Elton R, Van Tongeren M, Cohen GR et al. Interacting effects of particulate pollution and cold temperature on cardiorespiratory mortality in Scotland. Occup Environ Med 2008;65(3):197-204.

Dockery DW, Schwartz J, Spengler JD. Air pollution and daily mortality: associations with particulates and acid aerosols. Environ Res 1992;59(2):362-73.

Filleul L, Zeghnoun A, Cassadou S, Declercq C, Eilstein D, Le Tertre A, Medina S. et al. Influence of setup conditions of exposure indicators on the estimate of short-term associations between urban pollution and mortality. Science of the Total Environment 2006;355(1-3):90-97.

Fischer P, Ameling C, Marra M, Cassee FR. Absence of trends in relative risk estimates for the association between Black Smoke and daily mortality over a 34 years period in The Netherlands. Atmospheric Environment 2009;43(3):481-5.

Garcia-Aymerich J, Tobias A, Anto JM, Sunyer J. Air pollution and mortality in a cohort of patients with chronic obstructive pulmonary disease: a time series analysis. J Epidemiol Community Health 2000;54(1):73-4.

Goldberg MS, Burnett RT, Bailar JC, Brook J, Bonvalot Y, Tamblyn R et al. The association between daily mortality and ambient air particle pollution in Montreal, Quebec 1. Nonaccidental mortality. Environ Res 2001;86(1):12-25.

Goodman PG, Dockery DW, Clancy L. Cause-specific mortality and the extended effects of particulate pollution and temperature exposure. Environ Health Perspect 2004;112(2):179-85.

Gryparis A, Forsberg B, Katsouyanni K, Analitis A, Touloumi G, Schwartz J, Samoli E et al. Acute effects of ozone on mortality from the "Air pollution and health: A European approach" project. American Journal of Respiratory and Critical Care Medicine 2004;170(10):1080-1087.

Gwynn RC, Burnett RT, Thurston GD. A time-series analysis of acidic particulate matter and daily mortality and morbidity in the Buffalo, New York, region. Environ Health Perspect 2000;108(2):125-33.

Gwynn RC, Burnett RT, Thurston GD. A time-series analysis of acidic particulate matter and daily mortality and morbidity in the Buffalo, New York, region. Environ Health Perspect. 2000; 108(2): 125–133.

Hoek G, Brunekreef B, Verhoeff A, van Wijnen J, Fischer P. Daily mortality and air pollution in the Netherlands. J Air Waste Manage Assoc 2000;50(8):1380-1389.

Hong YC, Leem JH, Ha EH, Christiani DC. PM10 exposure, gaseous pollutants, and daily mortality in Inchon, South Korea. Environ Health Perspect 1999;107(11):873-888.

Katsouyanni K, Touloumi G, Spix C, Schwartz J, Balducci F, Medina S, Rossi G et al. Short-term effects of ambient sulphur dioxide and particulate matter on mortality in 12 European cities: results from time series data from the APHEA project. Air Pollution and Health: a European Approach. BMJ 1997;213(7095):1658-1663.

Katsouyanni K, Touloumi G, Samoli E, Gryparis A, Le Tertre A, Monopolis Y et al. Confounding and effect modification in the short-term effects of ambient particles on total mortality: Results from 29 European cities within the APHEA2 project. Epidemiol 2001;12(5):521-31.

Klemm RJ, Mason RM. Aerosol Research and Inhalation Epidemiological Study (ARIES): Air quality and daily mortality statistical modeling - Interim results. J Air Waste Manage Assoc 2000;50(8):1433-9.

Le Tertre A, Quenel P, Eilstein D, Medina S, Prouvost H, Pascal L et al. Short-term effects of air pollution on mortality in nine French cities: A quantitative summary. Arch Environ Health 2002;57(4):311-9.

Lippmann, M., Ito, K, Nadas, A., and Burnett, R. T. Association of particulate matter components with daily mortality and morbidity in urban populations. Health Effects Institute; 2000. Report No.: 95.

Maynard D, Coull BA, Gryparis A, Schwartz J. Mortality risk associated with short-term exposure to traffic particles and sulfates. Environ Health Perspect 2007;115(5):751-5.

Michelozzi P, Forastiere F, Fusco D, Perucci CA, Ostro B, Ancona C, Pallotti G. Air pollution and daily mortality in Rome, Italy. Occup Environ Med 1998;55(9):605-610.

Ostro B, Feng WY, Broadwin R, Green S, Lipsett M. The effects of components of fine particulate air pollution on mortality in California: Results from CALFINE. Environ Health Perspect 2007;115(1):13-9.

Parodi S, Vercelli M, Garrone E, Fontana V, Izzotti A. Ozone air pollution and daily mortality in Genoa, Italy between 1993 and 1996. Public Health 2005;199(9):844-850.

Pattenden S, Nikiforov B, Armstrong BG. Mortality and temperature in Sofia and London. J Epidemiol Community Health 2003;57(8):628-33.

Penttinen P, Tittanen P, Pekkanen J. Mortality and air pollution in metropolitan Helsinki, 1988-1996. Scandinavian Journal of Work Environment & Health 2004;30:19-27.

Perez Boillos MJ, Lopez AA, Estibalez Gonzalez JJ, Garcia Calabuig MA. The short-term effects of air pollution on mortality. The results of the EMECAM project in Vitoria-Gasteiz, 1990-94. Estudio Multicentrico Espanol sobre la Relacion entre la Contaminacion Atmosferica y la Mortalidad. [Spanish]. Rev Esp Salud Publica 1999;73(2):283-92.

Qian Z, He Q, Lin HM, Kong L, Liao D, Yang N, Bentley CM, Xu S. Short-term effects of gaseous pollutants on cause-specific mortality in Wuhan, China. J Air Waste Manag Assoc 2007;57(7):785-793.

Saez M, Figueiras A, Ballester F, Perez-Hoyos S, Ocana R, Tobias A. Comparing meta-analysis and ecological-longitudinal analysis in time-series studies. A case study of the effects of air pollution on mortality in three Spanish cities. J Epidemiol Community Health 2001;55(6):423-32.

Saez M, Ballester F, Barcelo MA, Perez-Hoyos S, Bellido J, Tenias JM, Ocana R et al. A combined analysis of the short-term effects of photochemical air pollutants on mortality within the EMECAM project. Environ Health Perspect 2002;110(3):221-228.

Schwartz J, Dockery DW, Neas LM. Is daily mortality associated specifically with fine particles? J Air Waste Manage Assoc 1996;46(10):927-39.

Simpson RW, Williams G, Petroeschevsky A, Morgan G, Rutherford S. Associations between outdoor air pollution and daily mortality in Brisbane, Australia. Arch Environ Health; 1997;52(6):442-454.

Stankovic A, Nikic,D, Nikolic M, Bogdanovic,D. Short-term effects of air pollution on cardiovascular mortality in elderly in Nis, Serbia. Central European Journal of Public Health 2007;15(3):95-98.

Wojtyniak B, Piekarska T. Short term effect of air pollution on mortality in Polish urban populations - What is different? J Epidemiol Community Health 1996;50(Suppl. 1):s36-s41.

Wong CM, Ma S, Hedley AJ, Lam TH. Effect of air pollution on daily mortality in Hong Kong. Environ Health Perspect 2001;109(4):335-340.

Zhang YH, Huang W, London SJ, Song GX, Chen GH, Jiang LL, Zhao NQ, Chen BH, Kan HD. Ozone and daily mortality in Shanghai, China. Environ Health Perspect 2006;114(8):1227-1232.

IV. American Cancer Society Cancer Prevention II Study (CPS II)

Cohort Information

Enrollment of the American Cancer Society CPS cohort occurred in 1982-1983 and was restricted to persons who were at least 30 years of age in households with at least one individual 45 years of age or older. After supplying written informed consent, subjects completed a questionnaire that included demographic characteristics, smoking history, alcohol use, diet and education¹. Mortality was ascertained by examination of death certificates in 1984, 1986, and 1988 and subsequently using the National Death Index². The study population included only those participants in CPS-II who resided in US Metropolitan Statistics Areas (MSAs) within the 48 contiguous states or the District of Columbia (based on their address at the time of enrollment), and for whom pollution data were available based on the presence of at least one pollution monitor for PM_{2.5}, sulfate, elemental carbon (EC) or ozone within their MSA.

Confounding variables

These variables included individual risk factors collected in the CPS-II questionnaire such as smoking habits and occupational status (see Jerrett et al. (2009)³ for a complete listing of the variables). Seven ecological covariates obtained from the 1980 US Census (median household income, proportion of the population <125% of the poverty line, percentage of unemployed persons over the age of 16 years, percentage of adults with education less than grade 12, percentage of homes with air conditioning, the Gini coefficient of income inequality and the percentage of the population that is white) were also included in the models. These variables were included at two levels: (1) as the MSA average and (2) as the zip code scale of residence for each subject, deviated from the MSA average (see Krewski et al. (2009)² for complete description of the variables).

Exposure Estimates for the American Cancer Society National Analysis

EC estimates were based on the Health Effects Institute (HEI) Air Quality Database (https://hei.aer.com/), which was accessed on June 18, 2009. Annual average values were already calculated based on accepted completeness criteria for the United States Environmental Protection Agency Speciation Trends Network. The completeness of data prior to 2000 limited the number of sites, and as a result we only extracted means for the years 2000-2007. Only sites with an HEI completeness value of 0.5 or above were included. After being joined to the metropolitan statistical areas (MSAs) or the primary metropolitan statistical areas (PMSA) through the longitude and latitude of the site, all monitoring sites within a given MSA/PMSA were averaged. Maximal spatial coverage was available for the years 2003-2006. A three-year average of the annual means was then calculated to assign exposure representing the long-term average concentration of EC in each MSA. In total we were able to join 66 MSAs with ACS health data and data on the other pollutants (PM_{2.5}, ozone, and sulfate). Details on the other pollutants and how they were calculated is available in Krewski et al. 2009⁴, Jerrett et al. 2009³, and Krewski et al. 2000⁵. These 66 MSAs formed the basis for our analysis presented in the main paper.

Statistical Modeling Details

We used Random Effects Cox (REC) models to estimate all the mortality risk estimates used in the American Cancer Society Cancer Prevention II National Analysis. The REC models were estimated based on custom software that we developed as part of an HEI project. Full details of the statistical model development and software are available in Krewski et al. (2009)⁴ and the software is also available publicly from HEI.

This model is expressed mathematically in the form:

$$h_{ik.s}(t) = h_{0s}(t) \eta_k * \exp(\beta' x_{ik})$$

where $h_{ik,s}(t)$ is the hazard function or instantaneous probability of death, at time t for the i^{th} subject in the k^{th} MSA in stratum s, conditional on the random effects η_k . Here, $h_{0s}(t)$ is the baseline hazard function for stratum s, defined by age-gender-race categories. The η_k are positive random variables representing the unexplained variation in the response between MSAs. Only the moments of the random effects need to be specified within our modeling framework:

$$E(\eta_i)=1$$
, $Var(\eta_k)=\tau^2$.

We assume that given the MSA random effects, responses between subjects are independent. Each model includes 20 individual variables expressed with 44 terms (for smoking, occupational exposures and the like) and the ecological variables as specified above.

Sources Cited: American Cancer Society Cancer Prevention Study II (CPS II)

- 1. Thun MJ, Calle EE, Namboodiri MM, al. e. Risk factors for fatal colon cancer in a large prospective study. J Natl Cancer Inst. 1992;84(491-500).
- 2. Calle EE, Terrell DD. Utility of the National Death Index for ascertainment of mortality among cancer prevention study II participants. Am J Epidemiol. 1993 Jan 15;137(2):235-41.
- 3. Jerrett M, Burnett RT, Pope CA, 3rd, Ito K, Thurston G, Krewski D, et al. Long-term ozone exposure and mortality. N Engl J Med. 2009 Mar 12;360(11):1085-95.
- 4. Krewski D, Jerrett M, Burnett RT, Ma R, Hughes E, Shi Y, et al. Extended analysis of the American Cancer Society study of particulate air pollution and mortality. Boston: Health Effects Institute; 2009.
- 5. Krewski D, Burnett R, Goldberg MS, Hoover K, Siemiatycki J, Jerrett M, et al. Reanalysis of the Harvard Six Cities Study and the American Cancer Society Study of Particulate Air Pollution and Mortality, Part II: Sensitivity Analysis: A Special Report of the Institute's Particle Epidemiology Reanalysis Project. Health Effects Institute, Cambridge, MA. Cambridge, MA; 2000